

Bayesian Statistics and R

Peng Ding,
School of Mathematical Sciences,
Peking Univ.

December 16, 2008

- What is *Probability*?
- Frequentist: LLN, CLT
- Bayesian: Prior and Posterior
- What is *Statistics*?

- Moment Estimation(Karl Pearson)
- Maximum Likelihood Estimation(Gauss, R. A. Fisher)
- Bayesian Method(Bayes)
- Empirical Bayesian(Robbins)
- ...?

Frequentist vs Bayesian

- Frequentist:
 - parameters are constant to be estimate
 - point estimation and interval estimation
- Bayesian:
 - parameters are random variables
 - Prior + Model \rightarrow Posterior
 - all information are contained in posterior distribution

Bayes' Formula and Bayesian Statistics

- Bayes' Formula:

If $\theta \in \Theta$ has prior distribution $\pi(\theta)$, and the observed data y comes from conditional distribution $p(y|\theta)$. Then the posterior distribution of θ given y is

$$\pi(\theta|y) = \frac{p(y|\theta)\pi(\theta)}{\int_{\Theta} p(y|\theta)\pi(\theta)d\theta}$$

- Bayesian Statistics:

- choose prior
- model observed data
- inference based on posterior distribution

How to choose *Prior*?

- Congugate Prior Distribution:
 - Prior and Posterior have the same form
 - Congugate Prior for Exponential Family

$$p(y_i|\theta) = f(y_i)g(\theta) \exp \left\{ \phi(\theta)^T u(y_i) \right\}$$

$$p(y|\theta) = \left(\prod_{i=1}^n f(y_i) \right) g(\theta)^n \exp \left\{ \phi(\theta)^T \sum_{i=1}^n u(y_i) \right\}$$

Choose Prior as:

$$p(\theta) \propto g(\theta)^\eta \exp\{\phi(\theta)^T \nu\}$$

The Posterior is:

$$p(\theta|y) \propto g(\theta)^{n+\eta} \exp\left\{ \phi(\theta)^T \left(\nu + \sum_{i=1}^n u(y_i) \right) \right\}$$

How to choose *Prior*?

- Non-informative Prior Distribution:
 - Bayesian Assumption

$$p(\theta) \propto \text{constant}, \theta \in \Theta$$

- Jefferys Prior

$$\pi(\theta) \propto |I(\theta)|^{1/2}$$

where $I(\theta)$ is the Fisher Information Matrix.

How to explore the *Posterior*?

- Direct Calculation:
 - The posterior has explicit and simple form!
- Simulation the Posterior:
 - Sampling from the posterior distribution
 - Markov Chain Monte Carlo(MCMC):
 - Gibbs Sampler and Metropolis- Hasting Algorithm

Gibbs Sampling

- $X \sim \pi(x), x = (x_1, \dots, x_n)$
- Initial value $x^{(0)} = (x_1^{(0)}, \dots, x_n^{(0)})$
- The t -th iteration:
 - Sample $x_1^{(t)} \sim \pi(x_1 | x_2^{(t-1)}, \dots, x_n^{(t-1)})$;
 - ...
 - Sample $x_i^{(t)} \sim \pi(x_i | x_1^{(t)}, \dots, x_{i-1}^{(t)}, x_{i+1}^{(t-1)}, \dots, x_n^{(t-1)})$;
 - ...
 - Sample $x_n^{(t)} \sim \pi(x_n | x_1^{(t)}, \dots, x_{n-1}^{(t)})$.
- Under some regular conditions, the distribution of x converges to the stationary distribution of the Markov Chain: $\pi(x)$.

M-H Algorithm

- Given an irreducible transition probability $q(\cdot, \cdot)$,
- Given a function

$$\alpha(\cdot, \cdot) = \min\left\{1, \frac{\pi(x')q(x', x)}{\pi(x)q(x, x')}\right\}, 0 < \alpha \leq 1.$$

- At time t , $X^{(t)} = x$,
 - Generate a potential transition $x \rightarrow x'$ by $q(x, \cdot)$;
 - With probability $\alpha(x, x')$, accept x' ; with probability $1 - \alpha(x, x')$ stay at x .
- Under some regular conditions, $\pi(x)$ is the stationary distribution of this Markov Chain.

- MCMCpack

```
library(lattice)
```

```
library(coda)
```

```
library(MASS)
```

```
library(MCMCpack)
```

- Famous software WinBUGS: Bayesian inference Using Gibbs Sampling.

Example 1: Binomial Distribution

- Model

$$p(y|\theta) \propto \theta^y(1-\theta)^{n-y}$$

- Prior

$$p(\theta) \propto \theta^{\alpha-1}(1-\theta)^{\beta-1}$$

- Posterior

$$p(\theta|y) \propto \theta^{\alpha+y-1}(1-\theta)^{\beta+n-y-1}$$

- R code

```
posterior <- MCbinomialbeta(y=3,n=12,alpha=1,beta=1,mc=5000)
summary(posterior)
plot(posterior)
```

Result for Binomial Distribution

1. Empirical mean and standard deviation for each variable, plus standard error of the mean:

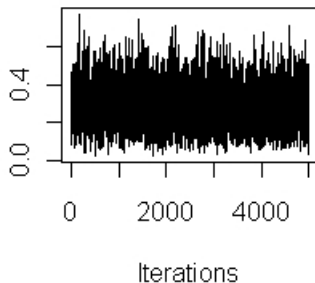
Mean	SD	Naive SE	Time-series SE
0.286442	0.116600	0.001649	0.001924

2. Quantiles for each variable:

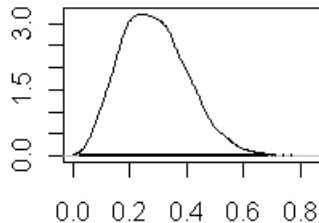
2.5%	25%	50%	75%	97.5%
0.08972	0.19926	0.27805	0.36167	0.53906

Posterior for Binomial Parameter

Trace of π



Density of π



N = 5000 Bandwidth = 0.0225

Example 2: Poisson Distribution

- Model

$$p(y|\lambda) \propto \prod_{i=1}^n \lambda^{y_i} e^{-\lambda}$$

- Prior

$$p(\lambda) \propto e^{\beta\lambda} \lambda^{\alpha-1}$$

- Posterior

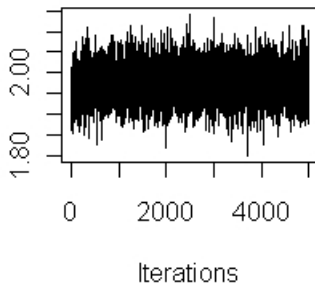
$$\lambda|y \sim \text{Gamma}(\alpha + n\bar{y}, \beta + n)$$

- R code

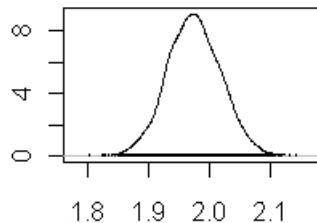
```
y<-rpois(1000,lambda=2)
posterior <- MCrpoissongamma(y, 15, 1, 5000)
summary(posterior)
plot(posterior)
```

Posterior for Poisson Parameter

Trace of lambda



Density of lambda



N = 5000 Bandwidth = 0.008447

Example 3: Normal Distribution with Variance known

- Model

$$p(y|\mu) \propto \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2\right\}$$

- Prior

$$p(\mu) \propto \exp\left\{-\frac{1}{2\tau_0^2} (\mu - \mu_0)^2\right\}$$

- Posterior

$$\mu|y \sim N(\mu_1, \tau_1^2)$$

where

$$\mu_1 = \frac{\frac{\mu_0}{\tau_0^2} + \frac{n\bar{y}}{\sigma^2}}{\frac{1}{\tau_0^2} + \frac{1}{\sigma^2}}, \bar{y} = \sum_{i=1}^n y_i/n,$$

$$\frac{1}{\tau_1^2} = \frac{1}{\tau_0^2} + \frac{1}{\sigma^2}.$$

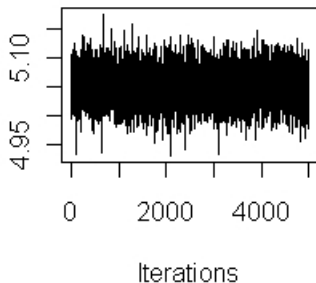
Example 3: Normal Distribution with Variance known

- R code

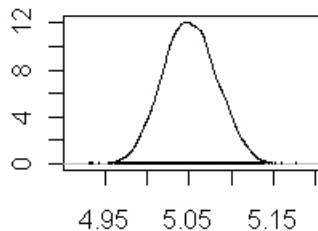
```
y<-rnorm(1000,5,1)
posterior <-
MCnormalnormal(y, sigma2=1, mu0=0,tau20=100, mc=5000)
summary(posterior)
plot(posterior)
```

Posterior for Normal Parameter: μ

Trace of mu



Density of mu



N = 5000 Bandwidth = 0.006083

Example 4: Normal Distribution with unknown Variance

- Model

$$p(y|\mu) \propto \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2\right\}$$

- Prior(Semi-Congugate)

$$\mu \sim N(b_0, B_0^2)$$

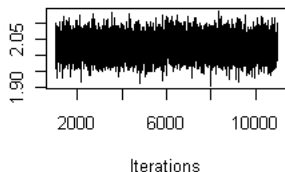
$$\sigma^2 \sim \text{Inverse} - \chi^2(2c_0, 2d_0)$$

- R code

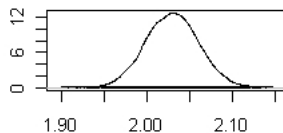
```
y<-rnorm(1000,2,1)
posterior<-
MCMCregress(y~1, b0 = 0, B0 =0, c0= 0.001, d0 = 0.001)
summary(posterior)
plot(posterior)
```

Posterior for Normal Parameter: μ and σ

Trace of (Intercept)

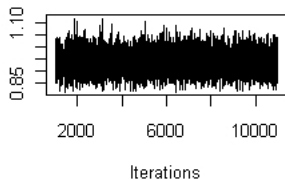


Density of (Intercept)

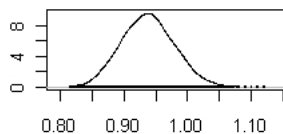


N = 10000 Bandwidth = 0.005113

Trace of sigma2



Density of sigma2



N = 10000 Bandwidth = 0.007073

Example 5: Multinomial Distribution

- Model

$$p(y|\theta) \propto \prod_{i=1}^n \theta_i^{y_i}$$

- Prior

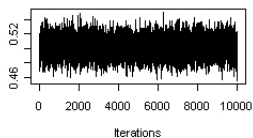
$$p(\theta|\alpha) \propto \prod_{i=1}^n \theta_i^{\alpha_i-1}$$

- R code

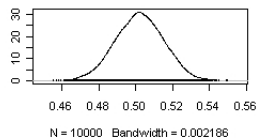
```
posterior <-  
MCMultinomialDirichlet(c(727,583,137), c(1,1,1), mc=10000)  
summary(posterior)  
plot(posterior)
```

Posterior for Multinomial Parameter

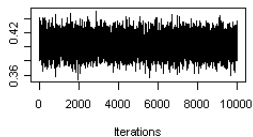
Trace of π_1



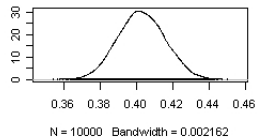
Density of π_1



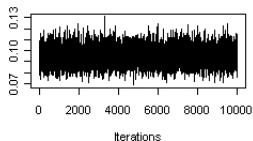
Trace of π_2



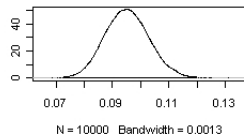
Density of π_2



Trace of π_3



Density of π_3



Generalized Linear Model

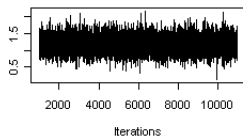
- $E(y|x) = g^{-1}(\beta^T x)$
- Different $g(\cdot)$ -Link Function, different models
 - Linear Regression: $g(t) = t: E(y|x) = \beta^T x$
 - Logistic Regression: $g(t) = \log\left(\frac{t}{1-t}\right): \log\left(\frac{P(y=1|x)}{1-P(y=1|x)}\right) = \beta^T x$
 - Probit Regression: $g(t) = \Phi^{-1}(t): \Phi^{-1}(P(y=1|x)) = \beta^T x$
 - Poisson Regression: $g(t) = \log(t): \log(E(y|x)) = \beta^T x$

Exmample 6: Linear Regression

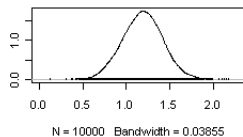
```
X<-rnorm(100,2,1)
Y<-1+2*X+rnorm(100,0,1)
posterior  <- MCMCregress(Y~X,b0 = 0, B0 = 0,
                          c0 = 0.001, d0 = 0.001,verbose=1000)
plot(posterior)
summary(posterior)
```

Posterior for Linear Regression

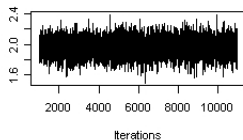
Trace of (Intercept)



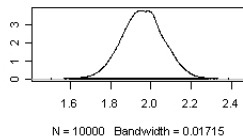
Density of (Intercept)



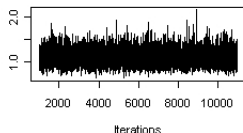
Trace of X



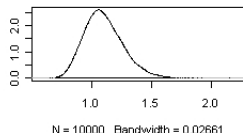
Density of X



Trace of sigma2



Density of sigma2

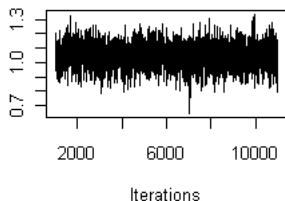


Example 7: Logistic Regression

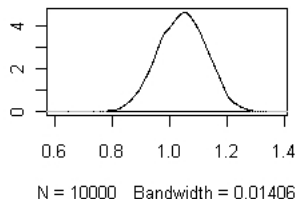
```
x<-rnorm(1000)
y<-rbinom(1000,1,exp(1-x)/(1+exp(1-x)))
posterior <-MCMClogit(y~x, b0=0, B0=.001)
plot(posterior)
summary(posterior)
```

Posterior for Logistic Regression

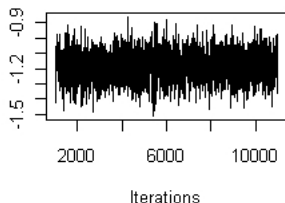
Trace of (Intercept)



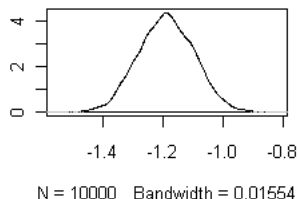
Density of (Intercept)



Trace of x



Density of x

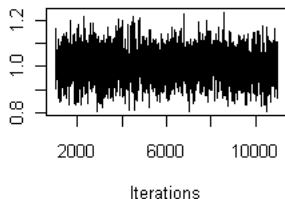


Example 8: Probit Regression

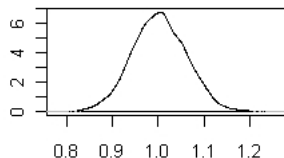
```
y<-rbinom(1000,1,pnorm(1-x))
posterior <- MCMCprobit(y~x, b0=0,B0=.001)
plot(posterior)
summary(posterior)
```

Posterior for Probit Regression

Trace of (Intercept)

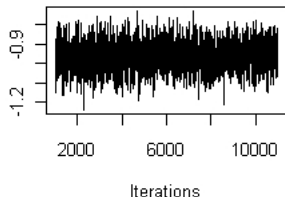


Density of (Intercept)

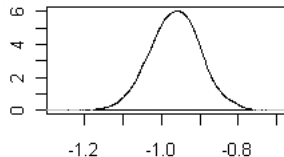


N = 10000 Bandwidth = 0.009996

Trace of x



Density of x



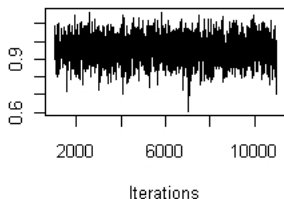
N = 10000 Bandwidth = 0.01089

Example 9: Poisson Regression

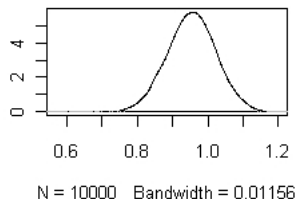
```
x<-rnorm(100)
y<-rpois(100,exp(1+x))
posterior <- MCMCpoisson(y ~x)
plot(posterior)
summary(posterior)
```

Posterior for Probit Regression

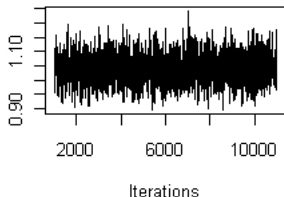
Trace of (Intercept)



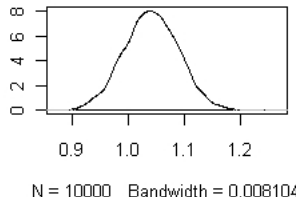
Density of (Intercept)



Trace of x



Density of x



- Gaussian Mixture Model
- Latent Class Analysis
- Hierarchical Models
- Perhaps any parametric models: Examples in WinBUGS
- Bayesians believes that all inference and more is Bayesian territory.—*Bayesian Nonparametrics*, J. K. Ghosh and R. V. Ramamoorthi, Springer(2003)

Reference and recommendatory books

- Andrew Gelman, John B Carlin, Hal S Stern and Donald B Rubin(2004), *Bayesian Data Analysis*, Chapman&Hall/CRC
- Martin A. Tanner(1996), *Tools for Statistical Inference: Methods for Exploration of Posterior Distribution and Likelihood Functions*, Springer
- Bradley P. Carlin and Thomas A. Liou(2000), *Bayes and Empirical Bayes Methods for Data Analysis*, Chapman&Hall/CRC
- Mao Shi-song, Wang Jing-long and Pu Xiao-long(2006), *Advanced Mathematical Statistics*, Higher Education Press(in Chinese).

Acknowledgement

- I would like to thank the organizers of the first R conference in China.
- Also I am grateful to my mentor Professor Zhi Geng for his introduction of GLM and Bayesian Methods.

Thank you!